Check for updates

PICRUSt2 for prediction of metagenome functions

To the Editor — One limitation of microbial community marker-gene sequencing is that it does not provide information about the functional composition of sampled communities. PICRUSt1 was developed in 2013 to predict the functional potential of a bacterial community on the basis of marker gene sequencing profiles, and now we present PICRUSt2 (https://github.com/ picrust/picrust2), which improves on the original method. Specifically, PICRUSt2 contains an updated and larger database of gene families and reference genomes, provides interoperability with any operational taxonomic unit (OTU)-picking or denoising algorithm, and enables phenotype predictions. Benchmarking shows that PICRUSt2 is more accurate than PICRUSt and other competing methods overall. PICRUSt2 also allows the addition of custom reference databases. We highlight these improvements and also important caveats regarding the use of predicted metagenomes.

The most common method for profiling bacterial communities is to sequence the conserved 16S rRNA gene. Functional profiles cannot be directly identified using 16S rRNA gene sequence data owing to strain variation, so several methods have been developed to predict microbial community functions from taxonomic profiles (amplicon sequences) alone¹⁻⁵. Shotgun metagenomics sequencing (MGS), which sequences entire genomes rather than marker genes, can also be used to characterize the functions of a community, but does not work well if there is host contamination — for example, in a biopsy — or if there is very little community biomass.

PICRUSt (hereafter "PICRUSt1") was developed for prediction of functions from 16S marker sequences, and it is widely used but has some limitations. Standard PICRUSt1 workflows require input sequences to be OTUs generated from closed-reference OTU-picking against a compatible version of the Greengenes database⁶. Due to this restriction to reference OTUs, the default PICRUSt1 workflow is incompatible with sequence denoising methods, which produce amplicon sequence variants (ASVs) rather than OTUs. ASVs have finer resolution, allowing closely related organisms to be more readily distinguished. Furthermore, the bacterial reference databases used by PICRUSt1



Fig. 1 | **PICRUSt2 algorithm. a**, The PICRUSt2 method consists of phylogenetic placement, hidden-state prediction and sample-wise gene and pathway abundance tabulation. ASV sequences and abundances are taken as input, and gene family and pathway abundances are output. All necessary reference tree and trait databases for the default workflow are included in the PICRUSt2 implementation. **b**, The default PICRUSt1 pipeline restricted predictions to reference OTUs in the Greengenes database. This requirement resulted in the exclusion of many study sequences across four representative 16S rRNA gene sequencing datasets. PICRUSt2 relaxes this requirement and is agnostic to whether the input sequences are within a reference database or not, which results in almost all of the input ASVs being retained in the final output. **c**, An increase in the taxonomic diversity in the default PICRUSt2 database is observed compared to PICRUSt1.

have not been updated since 2013 and lack thousands of recently added gene families.

We expected that optimizing genome prediction would improve accuracy of functional predictions. Therefore, the PICRUSt2 algorithm (Fig. 1a) includes steps that optimize genome prediction, including placing sequences into a reference phylogeny rather than relying on predictions limited to reference OTUs (Fig. 1b); basing predictions on a larger database of reference genomes and gene families (Fig. 1c); more stringently predicting pathway abundance (Supplementary Fig. 1); and enabling predictions of complex phenotypes and integration of custom databases.

PICRUSt2 integrates existing open-source tools to predict genomes of environmentally sampled 16S rRNA gene sequences. ASVs are placed into a reference tree, which is used as the basis of functional predictions. This reference tree contains 20,000 full 16S rRNA genes from bacterial and archaeal genomes in the Integrated Microbial Genomes (IMG) database⁷. Phylogenetic placement in PICRUSt2 is based on the output of three



Fig. 2 | **PICRUSt2 performance characteristics.** Validation of PICRUSt2 KO predictions comparing metagenome prediction performance against gold-standard shotgun MGS. **a**, Box plots of correlations observed in stool samples from Cameroonian individuals (n = 57), the Human Microbiome Project (HMP, n = 137), stool samples from Indian individuals (n = 91), non-human primate stool samples (n = 77), other mammalian stool (n = 8), ocean water (n = 6) and blueberry soil (n = 22) datasets. The significance of paired-sample, two-tailed Wilcoxon tests is indicated above each tested grouping (*P < 0.05; **P < 0.001; ns, not significant). **b**, Comparison of significantly differentially abundant KOs between predicted metagenomes and MGS. Precision, recall and F1 score are reported for each category compared to the MGS data. Precision corresponds to the proportion of significant KOs for that category also significant in the MGS data. Recall corresponds to the proportion of significant KOs in the MGS data also significant for that category. The F1 score is the harmonic mean of these metrics. The subsets of the four datasets compared are indicated above each panel (the Cameroonian parasite is *Entamoeba*). Wilcoxon tests were performed on the KO relative abundances after normalizing by the median number of universal single-copy genes per sample. Significance was defined at a false discovery rate <0.05. The Shuffled ASVs category corresponds to PICRUSt2 predictions with ASV labels shuffled per dataset. The Alt. MGS category corresponds to an alternative MGS processing pipeline with reads aligned to the KEGG database rather than the default HUMAnN2 pipeline.

tools: HMMER (http://www.hmmer.org)

to place ASVs, EPA-ng⁸ to determine the optimal position of these placed ASVs in a reference phylogeny, and GAPPA⁹ to output a new tree incorporating the ASV placements. This results in a phylogenetic tree containing both reference genomes and environmentally sampled organisms, which is used to predict individual gene-family copy numbers for each ASV. This procedure is rerun for each input dataset, allowing users to utilize custom reference databases as needed, including those that may be optimized for the study of specific microbial niches.

As in PICRUSt1, hidden-state prediction approaches are used in PICRUSt2 to infer the genomic content of sampled sequences. The castor R package¹⁰, which is substantially faster than the approach used in PICRUSt1, is used for core hidden-state prediction functions. As in PICRUSt1, ASVs are corrected by their 16S rRNA gene copy number and then multiplied by their functional predictions to produce a predicted metagenome. PICRUSt2 also provides the ASV contribution of each predicted function, allowing taxonomy-informed statistical analyses to be conducted. Lastly, pathway abundances are inferred on the basis of structured pathway mappings, which are more conservative than the 'bag-of-genes' approach used in PICRUSt1.

The PICRUSt2 default genome database is based on 41,926 bacterial and archaeal genomes from the IMG database⁷ (8 November 2017), which is a >20-fold

increase over the 2,011 IMG genomes used by PICRUSt1. Many of the additional genomes are from strains of the same species and have identical 16S rRNA genes. We de-replicated the identical 16S rRNA genes across these genomes, which resulted in 20,000 final 16S rRNA gene clusters. The taxonomic diversity of the PICRUSt2 reference database is greater than that of PICRUSt1 (Fig. 1c). The clearest increases in diversity are at the species and genus levels (5.3-fold and 2.2-fold increases, respectively), but all taxonomic levels are more diverse, including the phylum level, where the coverage increased from 39 to 64 phyla (a 1.6-fold increase).

PICRUSt2 predictions based on several gene family databases are supported by



Fig. 3 | **PICRUSt2** accurately predicts MetaCyc pathways and phenotypes for characterizing overall environments. **a**, Correlations between PICRUSt2-predicted pathway abundances and gold-standard MGS. Results are shown for each validation dataset: stool from Cameroonian individuals, the Human Microbiome Project (HMP), stool from Indian individuals, other mammalian stool, ocean water, non-human primate stool and blueberry soil. These results are limited to the 575 pathways that could potentially be identified by PICRUSt2 and HUMAnN2. **b**, Performance of binary phenotype predictions based on three metrics: F1 score, precision and recall. Each point corresponds to one of the 41 phenotypes tested. Predictions assessed here are based on holding out each genome individually, predicting the phenotypes for that held-out genome, and comparing the predicted and observed values. The null distribution in this case is based on randomizing the phenotypes across the reference genomes and comparing to the actual values, which results in the same output for all three metrics. The *P*-values of paired-sample, two-tailed Wilcoxon tests are indicated above each tested grouping (**P* < 0.05 and ***P* < 0.001). In **a** the *y* axis is truncated below 0.5 rather than 0 to better visualize small differences between categories. The sample sizes in **a** are 57 (Cameroonian), 137 (HMP), 91 (Indian), 8 (mammal), 6 (ocean), 77 (primate) and 22 (soil).

default, including Kyoto Encyclopedia of Genes and Genomes¹¹ (KEGG) orthologs (KOs) and Enzyme Commission numbers (EC numbers) (Supplementary Table 1). PICRUSt2 distinctly improves on PICRUSt1 by including gene families more recently added to the KEGG database. Specifically, the total number of KOs is 10,543 in PICRUSt2, as compared to 6,909 in PICRUSt1, a 1.5-fold increase.

We validated PICRUSt2 metagenome predictions using samples from seven published datasets generated using both 16S rRNA marker-gene and MGS. We used three human-associated microbiome datasets: 57 stool samples from Cameroonian individuals, 91 stool samples from Indian individuals, and 137 samples spanning the human body from the Human Microbiome Project. We also used four non-human-associated datasets, including 77 non-human primate stool samples, eight other mammalian stool samples, six ocean samples, and 22 bulk soil and blueberry rhizosphere samples. These datasets present a good variation of types of sequences and environments (Supplementary Table 2).

PICRUSt2 KO predictions from 16S rRNA marker gene data were produced for each dataset. We compared these predictions to KO relative abundances profiled from the corresponding MGS metagenomes, which served as a gold standard to evaluate prediction performance. We performed the same analyses with four alternative prediction pipelines: PICRUSt1, Piphillin², PanFP³ and Tax4Fun2^{4,5}. Spearman correlation coefficients (hereafter "correlations") were calculated for matching samples between the predicted KO abundance and MGS KO abundance tables after filtering all tables to the 6,220 KOs that could be output by all tested databases (Fig. 2). The correlation metric represents the similarity in rank ordering of KO abundances between the predicted and observed data. The correlations based on PICRUSt2 KO predictions ranged from a mean of 0.79 (s.d. = 0.028; primate stool) to 0.88 (s.d. = 0.019; Cameroonian stool dataset). For all seven datasets, PICRUSt2 predictions were either better than or comparable to the best prediction method (paired-sample, two-tailed Wilcoxon tests P < 0.05). Correlations based on PICRUSt2 predictions were substantially better for non-human-associated datasets. This result could indicate an advantage of phylogenetically based methods over non-phylogenetically based methods, such as Piphillin, for environments poorly represented by reference genomes.

Gene families regularly co-occur within genomes, so the use of correlations to assess gene-table similarity may be limited by the lack of independence of gene families within a sample (Supplementary Fig. 2). To address this dependency, we compared the observed correlations between paired MGS and predicted metagenomes to correlations between MGS functions and a null reference genome, comprised of the mean gene family

abundance across all reference genomes. For all datasets, PICRUSt2 metagenome tables were more similar to MGS values than the null (Fig. 2a). However, this increase over the null expectation is predominately driven by each dataset's predicted genome content (rather than that of individual samples). This is demonstrated by the fact that these correlations are actually only slightly significantly higher than those observed when ASV labels are shuffled within a dataset (Supplementary Fig. 3). The observed correlations for the shuffled ASVs ranged from a mean of 0.77 (s.d. = 0.196; primate stool) to 0.84 (s.d. = 0.178; blueberry rhizosphere). Biologically these results are consistent with several patterns. First, gene families are correlated in copy number across diverse taxa (as captured by the Null dataset). Second, these correlations are stronger within than between environments (as shown by the difference between the Null and Shuffled ASV results). Lastly, environment-to-environment differences tend to be larger than sample-to-sample differences within an environment (as shown by the differences between PICRUSt2 predictions and the Shuffled ASV results).

A complementary approach for validating metagenome predictions is to compare the results of differential abundance tests on 16S-predicted metagenomes to MGS data. A recent analysis of Piphillin suggested that this tool outperforms PICRUSt2 on the basis of this approach¹². We similarly performed this evaluation on the KO predictions for four validation datasets (Fig. 2b; see Supplementary Methods and Supplementary Results). Overall, PICRUSt2 displayed the highest F1 score, the harmonic mean of precision and recall, compared to other prediction methods (ranging from 0.46 to 0.59; mean = 0.51; s.d. = 0.06). However, all prediction tools displayed relatively low precision, the proportion of significant KOs that were also significant in the MGS data. In particular, precision ranged from 0.38 to 0.58 (mean = 0.48; s.d. = 0.08) for PICRUSt2 and 0.06 to 0.66 (mean = 0.45; s.d. = 0.27) for Piphillin. In all cases, PICRUSt2 predictions outperformed ASV-shuffled predictions, which ranged in precision from 0.22 to 0.42 (mean = 0.30; s.d. = 0.09). In addition, differential abundance tests performed on MGS-derived KOs from an alternative MGS-processing workflow resulted in only marginally higher precision (ranging from 0.57 to 0.67; mean = 0.62; s.d. = 0.04). Taken together, these results highlight the difficulty of reproducing microbial functional biomarkers with both predicted and actual metagenomics data.

MetaCyc pathway abundances are now the main high-level predictions output by PICRUSt2 by default. The MetaCyc database¹³ is an open-source alternative to KEGG and is also a major focus of the widely used metagenomics functional profiler HUMAnN2¹⁴. MetaCyc pathway abundances are calculated in PICRUSt2 through structured mappings of EC gene families to pathways. These pathway predictions performed better than the null distribution for all metrics overall (paired-sample, two-tailed Wilcoxon tests P < 0.05; Fig. 3a and Supplementary Figs. 4 and 5) when compared to MGS-derived pathways. As in our previous analysis, shuffled ASV predictions representing overall functional structure within each dataset accounted for the majority of this signal (Supplementary Fig. 4). In addition, differential abundance tests on these pathways showed high variability in F1 scores across datasets and statistical methods, with the ASV shuffled predictions contributing the majority of this signal (Supplementary Fig. 6). F1 scores ranged from 0.23 to 0.62 (mean = 0.41; s.d. = 0.17) and 0.22 to 0.60 (mean = 0.34; s.d. = 0.18) for the observed and ASV shuffled PICRUSt2 predictions. respectively. Again, these results suggest that identifying robust differentially abundant metagenome-wide pathways is difficult, highlighting the challenge of analyzing microbial pathways in general.

Predictions for 41 microbial phenotypes, which are linked to IMG genomes15, can also now be generated with PICRUSt2. These represent high-level microbial metabolic activities such as "glucose utilizing" and "denitrifier" that are annotated as present or absent within each reference genome. We performed a hold-out validation to assess the performance of PICRUSt2 phenotype predictions, which involved comparing the binary phenotype predictions to the expected phenotypes for each reference genome. As based on F1 score (mean = 84.8%; s.d. = 9.01%), precision (mean = 86.5%; s.d. = 6.21%), and recall (mean = 83.5%; s.d. = 11.4%), these predictions performed significantly better than the null expectation (Fig. 3b; Wilcoxon tests P < 0.05).

There are two main criticisms of amplicon-based functional prediction. The first is that the predictions are biased toward existing reference genomes, which means that rare environment-specific functions are less likely to be identified. This limitation is decreasing over time as the number of high-quality available genomes continues to grow. PICRUSt2 also allows user-specified genomes to be used for generating predictions, which provides a flexible framework for studying particular environments. The second criticism is that amplicon-based predictions cannot provide resolution to distinguish strain-specific functionality. This is an important limitation of PICRUSt2 and any amplicon-based analysis, which can only differentiate taxa to the degree they differ at the amplified marker gene sequence.

PICRUSt2 provides improved accuracy and flexibility for marker gene metagenome inference. We have highlighted these improvements while also describing limitations in identifying consistent differentially abundant functions in microbiome studies. We hope that the expanded functionality of PICRUSt2 will continue to enable the identification of insights into functional microbial ecology from amplicon sequencing profiles.

Data availability

The repository at https://github.com/ gavinmdouglas/picrust2_manuscript includes the processed data files that can be used to re-generate the figures and findings in this paper. The accession codes for all sequencing data used in this study are listed in the Supplementary Methods.

Code availability

PICRUSt2 is available at https://github. com/picrust/picrust2. The Python and R code used for the analyses and database construction described in this paper are available online at https://github.com/ gavinmdouglas/picrust2 manuscript.

Editorial note: This article has been peer reviewed.

Gavin M. Douglas¹, Vincent J. Maffei^{D²}, Jesse R. Zaneveld³, Svetlana N. Yurgel⁴, James R. Brown 5, Christopher M. Taylor 2, Curtis Huttenhower b⁶ and Morgan G. I. Langille ^{1,7}

¹Department of Microbiology and Immunology, Dalhousie University, Halifax, Nova Scotia, Canada. ²Department of Microbiology, Immunology, and Parasitology, Louisiana State University Health Sciences Center, New Orleans, LA, USA. ³Division of Biological Sciences, School of STEM, University of Washington Bothell, Bothell, WA, USA. ⁴Department of Plant, Food, and Environmental Sciences, Dalhousie University, Truro, Nova Scotia, Canada. ⁵Computational Biology, GlaxoSmithKline R&D, Collegeville, PA, USA. Biostatistics Department, Harvard T. H. Chan School of Public Health, Boston, MA, USA. 7Department of Pharmacology, Dalhousie University, Halifax, Nova Scotia, Canada.

[™]e-mail: morgan.langille@dal.ca

Published online: 1 June 2020 https://doi.org/10.1038/s41587-020-0548-6

References

- 1. Langille, M. G. I. et al. Nat. Biotechnol. 31, 814-821 (2013).
- 2. Iwai, S. et al. PLoS One 11, e0166104 (2016).
- 3. Jun, S. R. et al. BMC Res. Notes 8, 479 (2015).
- Aßhauer, K. P., Wemheuer, B., Daniel, R. & Meinicke, P. 4. Bioinformatics 31, 2882-2884 (2015).
- Wemheuer, F. et al. Preprint at bioRxiv https://doi. 5. org/10.1101/490037 (2018).
- 6. DeSantis, T. Z. et al. Appl. Environ. Microbiol. 72, 5069-5072 (2006).
- 7. Markowitz, V. M. et al. Nucleic Acids Res. 40 (D1), D115-D122 (2012).
- Barbera, P. et al. Syst. Biol. 68, 365-369 (2019). 8.
- 9 Czech I. & Stamatakis A PLoS One 14 e0217050 (2019)
- 10. Louca, S. & Doebeli, M. Bioinformatics 34, 1053-1055 (2018).
- 11. Kanehisa, M., Goto, S., Sato, Y., Furumichi, M. & Tanabe, M.
- Nucleic Acids Res. 40 (D1), D109-D114 (2012). 12. Naravan, N. R. et al. BMC Genomics 21, 56 (2020).
- 13. Caspi, R. et al. Nucleic Acids Res. 44 (D1), D471-D480 (2016). 14. Franzosa, E. A. et al. Nat. Methods 15, 962-968 (2018).
- 15. Chen, I.-M. A. et al. PLoS One 8, e54859 (2013).

Acknowledgements

We thank Zhenjiang Xu and Amy Chen for providing us access to data files used for testing and the default reference database. We also thank Heather McIntosh for help designing the pipeline flowchart. G.M.D. is funded by an NSERC CGS-D scholarship. V.J.M. is funded by an NIH/NIAAA Ruth L. Kirschstein National Research Service Award (F30 AA026527). J.R.Z. is supported by NSF IOS CAREER grant 1942647. S.Y.N. is funded by an NSERC Discovery Grant. C.H. is funded in part by NIH NIDDK grants U54DK102557 and R24DK110499. M.G.I.L. is funded by an NSERC Discovery Grant and an NSERC Collaborative Research Development with co-funding from GlaxoSmithKline to M.G.I.L. and J.R.B.

Competing interests

The authors declare no competing interests.

Additional information

Supplementary information is available for this paper at https://doi.org/10.1038/s41587-020-0548-6.